

Martingales + Concentration of Measure:

Concentration: general phenomenon that if X_1, \dots, X_n only "weakly dependent", and $f(X_1, \dots, X_n)$ "weakly sensitive" to each input, then

$P[|f(X_1, \dots, X_n) - \mathbb{E}f(X_1, \dots, X_n)| > t]$ is "small".

One of main classical examples:

Thm: (Hoeffding neg.) X_1, \dots, X_n independent, $X_i \in [a_i, b_i]$ a.s. $\forall i$, $S_n := \sum_{i=1}^n X_i$.
(i.e. $f(X_1, \dots, X_n) = \sum_{i=1}^n X_i$). Then,

$P[|S_n - \mathbb{E}S_n| > t] \leq 2 \exp\left(-\frac{t^2}{2\sigma^2}\right)$, $\sigma^2 = \frac{1}{4} \sum_{i=1}^n (b_i - a_i)^2$.
 \approx tail of $N(0, \sigma^2)$

Ex: (SRW) $X_i \stackrel{iid}{\sim} \text{Unif}(\pm 1)$ $\rightarrow \sigma^2 = n$, $\mathbb{E}S_n = 0$, can rewrite Hoeffding:
 $P\left[\left|\frac{1}{\sqrt{n}} S_n\right| > t\right] \leq 2 \exp\left(-\frac{t^2}{2}\right) \rightarrow$ captures $S_n/\sqrt{n} \approx N(0, 1)$ (in tails)
 $\text{Var}(S_n/\sqrt{n}) = 1$

Proof technique: "Chernoff method": exponential Markov neg.
 $P[S \geq t] = P[\exp(\lambda S) \geq \exp(\lambda t)] \leq \frac{\mathbb{E} \exp(\lambda S)}{\exp(\lambda t)} = \frac{\phi_S(\lambda)}{\exp(\lambda t)}$

Def: $\phi_X(\lambda) := \mathbb{E} \exp(\lambda X)$ (moment generating function) = $\exp(\psi_X(\lambda) - \lambda t)$
 $\psi_X(\lambda) := \log \phi_X(\lambda)$ (cumulant generating function)

Prop: X, Y indep $\Rightarrow \phi_{X+Y}(\lambda) = \phi_X(\lambda) \phi_Y(\lambda)$, $\psi_{X+Y}(\lambda) = \psi_X(\lambda) + \psi_Y(\lambda)$.

Def: X is σ^2 -subgaussian if $\forall \lambda \in \mathbb{R}$, $\phi_{X-\mathbb{E}X}(\lambda) \leq \exp\left(\frac{\sigma^2 \lambda^2}{2}\right)$ } equivalent.

Prop: $X \sim N(0, \sigma^2) \Rightarrow \phi_X(\lambda) = \exp\left(\frac{\sigma^2 \lambda^2}{2}\right)$. $\psi_{X-\mathbb{E}X}(\lambda) \leq \frac{\sigma^2 \lambda^2}{2}$

Prop: X, Y indep., X σ^2 -subgauss., Y τ^2 -subgauss. $\Rightarrow X+Y$ is $(\sigma^2 + \tau^2)$ -subgauss.

Lem: If X σ^2 -subgaussian $\Rightarrow P[|X - \mathbb{E}X| > t] \leq 2 \exp\left(-\frac{t^2}{2\sigma^2}\right)$.

Pf: Chernoff method:
 $P[X - \mathbb{E}X > t] \leq \frac{\phi_{X-\mathbb{E}X}(\lambda)}{\exp(\lambda t)} \leq \exp\left(\frac{\sigma^2 \lambda^2}{2} - \lambda t\right)$ (quadratic in λ)
Optimal $\lambda^* = \frac{t}{\sigma^2} \rightarrow \leq \exp\left(-\frac{t^2}{2\sigma^2}\right)$

Remaining: show that $X \in [a, b]$ then X σ^2 -subgaussian.

Lem: (Hoeffding) If $X \in [a, b]$ a.s., then X is $\frac{1}{4}(b-a)^2$ -subgaussian. (VARIANCE PROXY)

Prop: If $X \in [a, b]$ a.s., then $\text{Var}(X) \leq \frac{1}{4}(b-a)^2$.

Pf: $\text{Var}(X) = \text{Var}\left(X - \frac{a+b}{2}\right) \leq \mathbb{E}\left(X - \frac{a+b}{2}\right)^2 \leq \left(\frac{b-a}{2}\right)^2 = \frac{1}{4}(b-a)^2$.

Prop: If $\psi_{X-\mathbb{E}X}(\lambda)$ smooth, $\psi_{X-\mathbb{E}X}''(\lambda) \leq \sigma^2 \forall \lambda \in \mathbb{R} \Rightarrow X$ σ^2 -subgaussian.

Pf: Idea: Taylor @ $\lambda=0$. Recall: $\psi_{X-\mathbb{E}X}(\lambda) = \log \mathbb{E} \exp(\lambda(X-\mathbb{E}X))$.

$\psi_{X-\mathbb{E}X}(0) = \log(1) = 0$. $\psi_{X-\mathbb{E}X}'(0) = \frac{\mathbb{E}(X-\mathbb{E}X) \exp(\lambda(X-\mathbb{E}X))}{\mathbb{E} \exp(\lambda(X-\mathbb{E}X))} \Big|_{\lambda=0} = 0$.

Taylor (w/ Lagrange remainder) $\Rightarrow \forall \lambda \exists \theta$ s.t. $\psi_{X-\mathbb{E}X}(\lambda) = \psi(0) + \lambda \psi'(0) + \frac{\lambda^2}{2} \psi''(\theta) \leq \frac{\sigma^2 \lambda^2}{2}$.

Pf: (of H. Lem) WLOG $\mathbb{E}X = 0$. Compute:

$\psi_X'(\lambda) = \frac{\mathbb{E} X \exp(\lambda X)}{\mathbb{E} \exp(\lambda X)}$, $\psi_X''(\lambda) = \frac{\mathbb{E} X^2 \exp(\lambda X)}{\mathbb{E} \exp(\lambda X)} - \left(\frac{\mathbb{E} X \exp(\lambda X)}{\mathbb{E} \exp(\lambda X)}\right)^2$

$\rho(x) := \frac{\exp(\lambda x)}{\mathbb{E} \exp(\lambda X)} = \mathbb{E} X \rho(X) = \mathbb{E} Y$ (fitting or "reweighting" of X)
 $= \mathbb{E} X^2 \rho(X) - (\mathbb{E} X \rho(X))^2 = \text{Var}(Y)$

$\psi_X''(\lambda) = \text{Var}(Y) \leq \frac{1}{4}(b-a)^2$.
i.e. $\rho \geq 0$, $\int \rho d\mu = 1 = \mathbb{E}_\rho(X)$. $Y \in [a, b]$ a.s.

Pf: (of H. neg.) $X_i \in [a_i, b_i]$ i.s. $\frac{1}{4}(b_i - a_i)^2$ -subgauss $\Rightarrow S_n = \sum X_i$ is $\left(\sum \frac{1}{4}(b_i - a_i)^2\right)$ -subgaussian.

Generalization to martingales:

Thm: (Azuma) $(M_n)_{n \geq 0}$ martingale, $(A_n), (B_n)$ predictable, $c_n > 0$ s.t.:

$A_i \leq M_i - M_{i-1} \leq B_i$ and $B_i - A_i \leq c_i$ a.s. (w/lt filtration $(\mathcal{F}_n)_{n \geq 0}$)

Then, $M_n - M_0$ is σ^2 -subgauss, $\sigma^2 = \frac{1}{4} \sum_{i=1}^n c_i^2$.

Pf: Tower rule:
 $\mathbb{E} \exp\left(\lambda \sum_{i=1}^n (M_i - M_{i-1})\right) = \mathbb{E} \left[\mathbb{E} \left[\exp\left(\lambda \sum_{i=1}^n (M_i - M_{i-1})\right) \mid \mathcal{F}_{n-1} \right] \right]$
 $= \mathbb{E} \left[\exp\left(\lambda \sum_{i=1}^{n-1} (M_i - M_{i-1})\right) \mathbb{E} \left[\exp\left(\lambda (M_n - M_{n-1})\right) \mid \mathcal{F}_{n-1} \right] \right]$
 $\leq \dots \leq \exp\left(\frac{\lambda^2}{2} \cdot \frac{c_n^2}{4}\right) \mathbb{E} \left[\exp\left(\lambda (M_{n-1} - M_0)\right) \right]$
 $\leq \dots$ repeat $(n-1)$ times $\leq \exp\left(\frac{\lambda^2}{2} \cdot \sigma^2\right)$.

Ex: Enough to have $|M_i - M_{i-1}| \leq b_i$ a.s. $\rightarrow A_i = -b_i, B_i = b_i \rightarrow c_i = 2b_i$
 $\rightarrow \sigma^2 = \sum b_i^2$.

Concrete version of general concentration principle:

Def: $f: \mathcal{X}_1 \times \dots \times \mathcal{X}_n \rightarrow \mathbb{R}$, $\delta_i(f) := \sup \{ |f(x) - f(x')| : x, x' \text{ only differ in coord } i \}$.

Cor: (Slightly weak McDiarmid) X_1, \dots, X_n any indep. r.v. (not necessarily $\in \mathbb{R}$)
 f is s.t. $\delta_i(f) < \infty$, $\mathbb{E} |f(X_1, \dots, X_n)| < \infty$, then

$f(X_1, \dots, X_n)$ is σ^2 -subgauss, where $\sigma^2 = \sum_{i=1}^n \delta_i(f)^2$.

$P[|f(X_1, \dots, X_n) - \mathbb{E}f(X_1, \dots, X_n)| > t] \leq 2 \exp\left(-\frac{t^2}{2\sigma^2}\right)$.

Pf: Use Doob martingale: $M_i := \mathbb{E}[f(X_1, \dots, X_n) \mid X_1, \dots, X_i]$
(i.e., $\mathcal{F}_i = \sigma(X_1, \dots, X_i)$)

$M_0 = \mathbb{E}f(X_1, \dots, X_n)$, $M_n = f(X_1, \dots, X_n)$
 $\Rightarrow M_n - M_0 = f - \mathbb{E}f$. (Rk: Law(X) = Law(X'))

Just need: $|M_i - M_{i-1}| \leq \delta_i(f)$ a.s.
 (X_1^i, \dots, X_n^i) ind. copy of (X_1, \dots, X_n) ,
 $X^{(i)} = (X_1, \dots, X_{i-1}, X_i^i, X_{i+1}, \dots, X_n)$. (resample coord i of X)

$M_i - M_{i-1} = \mathbb{E}[f(X) \mid X_1, \dots, X_i] - \mathbb{E}[f(X) \mid X_1, \dots, X_{i-1}]$
 $= \mathbb{E}[f(X) - f(X^{(i)}) \mid X_1, \dots, X_i] = \mathbb{E}[f(X^{(i)}) \mid X_1, \dots, X_{i-1}]$
 $= \mathbb{E}[f(X^{(i)}) \mid X_1, \dots, X_i]$
 $\in [-\delta_i(f), \delta_i(f)]$.

Use Azuma w/ $c_i = 2\delta_i = 2\delta_i(f)$.

Ex: (Balls + bins (coupon collector)) throw m balls into n bins unit at random independently. $Z := \#\{\text{empty bins}\} = \sum_{i=1}^n \mathbb{1}\{\text{bin } i \text{ empty}\}$

$\mathbb{E}Z = n P[\text{bin } 1 \text{ empty}] = n \left(\frac{n-1}{n}\right)^m = n \left(1 - \frac{1}{n}\right)^m$
e.g. $n = cm$, $m \rightarrow \infty$, $\mathbb{E}Z \sim cm \exp\left(-\frac{1}{c}\right) = \frac{c}{e^{1/c}} \cdot (m)$.

$X_j =$ index of bin ball j lands in ($\in \{1, \dots, n\}$, for $j = 1, \dots, m$).
 $Z = f(X_1, \dots, X_m)$ $\delta_j(f) = 1 \rightarrow \sigma^2 = m$.

$\Rightarrow P[|Z - \mathbb{E}Z| \geq t] \leq 2 \exp\left(-\frac{t^2}{2m}\right)$. (Rk: No model-specific thinking!)
e.g. w/prob ≥ 0.999 , $|Z - \mathbb{E}Z| \leq C\sqrt{m}$.
 $\Theta(\sqrt{m})$